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"Probability is good sense reduced to a calculus" (S. Laplace)













Two-photon invariant mass

ATLAS Experiment at LHC (CERN, Geneva)





ATLAS Experiment at LHC [length: 46 m; Ø 25 m]



pprox 7000 tonnes

pprox 100 millions electronic channels



Two flashes of 'light' (2 γ 's) in a 'noisy' environment.



Two flashes of 'light' (2 γ 's) in a 'noisy' environment. Higgs $\rightarrow \gamma \gamma$?



Two flashes of 'light' (2 γ 's) in a 'noisy' environment. Higgs $\rightarrow \gamma \gamma$? Probably not...







Quite indirect measurements of something we do not "see"!

But, can we see our mass?





... or a voltage?





... or our blood pressure?





Certainly not!

Certainly not!

- ... although for some quantities we can have
- a 'vivid impression' (in the David Hume's sense)



Measuring a mass on a scale



Equilibrium:

 $mg - k\Delta x = 0$ $\Delta x \rightarrow \theta \rightarrow \text{scale reading}$

(with 'g' gravitational acceleration; 'k' spring constant.)



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(with 'g' gravitational acceleration; 'k' spring constant.)

From the reading to the value of the mass:

scale reading $\xrightarrow{given g, k, "etc."...} m$

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scale reading
$$\xrightarrow{given g, k, "etc."...} m$$

Dependence on 'g': $g \stackrel{?}{=} \frac{GM_{t}}{R_{t}^2}$









... not even ellipsoidal...





- Earth not spherical...
- ... not even ellipsoidal...
- ...and not even homogeneous.



- Position is usually <u>not</u> at "R_b" from the Earth center;
- Earth not spherical...
- ... not even ellipsoidal...
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- Moreover we have to consider centrifugal effects



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▶





left to your imagination...


Measuring a mass on a balance



 $\Delta x \rightarrow \theta \rightarrow$ scale reading:

left to your imagination...

- + randomic effects:
 - stopping position of damped oscillation;
 - variability of all quantities of influence (in the ISO-GUM sense);
 - reading of analog scale.

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$\mathsf{Mass} \longrightarrow \mathsf{Reading}$



$\mathsf{Mass} \longrightarrow \mathsf{Reading}$





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$\mathsf{Reading} \longrightarrow `\mathsf{true'} \ \mathsf{mass}$



1 incomplete definition of the measurand



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 $\rightarrow g \\ \rightarrow where? \\ \rightarrow inertial effects subtracted?$



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ightarrow gightarrowinertial effects subtracted?

2 imperfect realization of the definition of the measurand



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- \rightarrow scattering on neutron
 - $\rightarrow \mathrm{how}$ to realize a neutron target?



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- 4 inadequate knowledge of the effects of environmental conditions on the measurement, or imperfect measurement of environmental conditions;
- 5 personal bias in reading analogue instruments;

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Note

- Sources not necessarily independent
- In particular, sources 1-9 may contribute to 10 (e.g. not-monitored electric fluctuations)

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Error and uncertainty are not synonyms!

Observation \rightarrow value of a quantity



scale reading
$$\xrightarrow{given g, k, "etc."...} m$$

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$Observations \rightarrow hypotheses$

This problem occurs not only "determining" *the* value of a physical quantity.



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• Experimental observation ('data') \rightarrow responsible cause.



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• Experimental observation ('data') \rightarrow responsible cause.

(But logically no substantial difference.)



$\mathsf{Causes} \to \mathsf{effects}$

The same apparent cause might produce several, different effects



Given an observed effect, we are not sure about the exact cause that has produced it.

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 $\mathsf{E}_2 \Rightarrow \{\mathit{C}_1, \ \mathit{C}_2, \ \mathit{C}_3\}?$



The "essential problem" of the Sciences

"Now, these problems are classified as *probability of causes*, and are most interesting of all for their scientific applications. I play at *écarté* with a gentleman whom I know to be perfectly honest. What is the chance that he turns up the king? It is 1/8. This is a problem of the probability of effects.


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I play with a gentleman whom I do not know. He has dealt ten times, and he has turned the king up six times. What is the chance that he is a sharper? This is a problem in the probability of causes. It may be said that it is the essential problem of the experimental method."

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Why we (or most of us) have not been taught how to tackle this kind of problems?

An example easy to understand:

- two causes;
- two effects;



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- medical diagnostics helps to clarify the issues:
 - easier to reach intuitive answers

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An example easy to understand:

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medical diagnostics helps to clarify the issues:

- easier to reach intuitive answers
- ...although if someone might have fallacious intuitions
 - \Rightarrow a formal guide helps us avoiding errors
 - \Rightarrow logics of the uncertain (theory of probabilities)

An Italian citizen is selected at random to undergo an AIDS test.

 \rightarrow Performance of clinical trial is not perfect, as customary:

P(Pos | HIV) = 100% $P(\text{Pos} | \overline{\text{HIV}}) = 0.2\%$ $P(\text{Neg} | \overline{\text{HIV}}) = 99.8\%$ $H_1 = \text{'HIV'} \text{ (Infected)} \qquad E_1 = \text{Positive}$ $H_2 = \text{'HIV'} \text{ (Not infected)} \qquad E_2 = \text{Negative}$

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Result: \Rightarrow <u>Positive</u>

Infected or not infected?

Being $P(Pos | \overline{HIV}) = 0.2\%$ and having observed 'Positive', can we say?

"It is practically impossible that the person is not infected, since it was practically impossible that a non infected person would result positive"



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(We will learn in the sequel how to evaluate it correctly)

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Instead, $P(\text{HIV} | \text{Pos, random Italian}) \approx 45\%$

 $\Rightarrow Serious mistake! (not just 99.8\% instead of 98.3\% or so)$... from which bad decisions might follow!

???

Where is the problem?



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The previous statements, although dealing with probabilistic issues, **are not ground** on probability theory



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- ... and in these issues intuition can be fallacious!
- \Rightarrow A sound formal guidance can rescue us



Pay attention not to arbitrary revert conditional probabilities:

In general $P(A | B) \neq P(B | A)$



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$\begin{array}{rl} P(E \mid H) \lll 1 & \underline{\text{does not imply}} & P(H \mid E) \lll 1 \\ & & (\text{ and '}\underline{\text{hence'}} & P(\overline{H} \mid E) \approx 1 \,) \end{array}$

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 \Rightarrow Prosecutor's fallacy

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 \Rightarrow Misunderstanding **p-values** (a *related* logical mistake)

\rightarrow Probability of causes

"the essential problem of the experimental method"



From 'true value' to observations



Given μ (exactly known) we are uncertain about x



From 'true value' to observations



Uncertainty about μ makes us more uncertain about x




The observed data is certain: \rightarrow 'true value' uncertain.

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The observed data is certain: \rightarrow 'true value' uncertain. "data uncertainty" ?



The observed data is certain: \rightarrow 'true value' uncertain. "data uncertainty" ? Data corrupted?



The observed data is certain: \rightarrow 'true value' uncertain.

"data uncertainty" ? Data corrupted? Even if the data were corrupted, the <u>data</u> were the corrupted data!!...



Where does the observed value of x comes from?

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We are now uncertain about μ , given x.

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Note the symmetry in reasoning.

Basic rules of probability

1.
$$0 \leq P(A \mid I) \leq 1$$

$$2. \quad P(\Omega \mid \mathbf{I}) = 1$$

3.
$$P(A \cup B \mid I) = P(A \mid I) + P(B \mid I)$$
 [if $P(A \cap B \mid I) = \emptyset$]

4.
$$P(A \cap B | I) = P(A | B, I) \cdot P(B | I) = P(B | A, I) \cdot P(A | I)$$

Remember that probability is always conditional probability!

/ is the background condition (related to information ${}^{\prime}I'_{s}$) \rightarrow usually implicit (we only care about 're-conditioning')

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- $\rightarrow\,$ usually implicit (we only care about 're-conditioning')
- Note: 4. <u>does not</u> define conditional probability. (Probability is <u>always</u> conditional probability!)
 - \Rightarrow easily extended to uncertain numbers ('random variables')

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(Schrödinger, 1947)



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Probability depends on the status of information of the *subject* who evaluates it.



"Thus whenever we speak loosely of 'the probability of an event', it is always to be understood: probability with regard to a certain given state of knowledge"

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$P(E) \longrightarrow P(E \mid I_s(t))$

where $I_s(t)$ is the information available to subject s at time t.

Mathematics of beliefs

An even better news:

The fourth basic rule can be fully exploited!



Mathematics of beliefs

An even better news:

The fourth basic rule can be fully exploited!

(Liberated by a curious ideology that forbids its use)





$P(A \mid B \mid I) P(B \mid I) = P(B \mid A, I) P(A \mid I)$

 $P(A|B) = \frac{P(B|A) P(A)}{P(A)}$

 \bigcirc

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Take the courage to use it!

 $P(A|B) = \frac{P(B|A) P(A)}{P(B)}$

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$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$ It's easy if you try...! © GdA, LTP/PSI 05/05/22, 29/71



"The greater the probability of an observed event given any one of a number of causes to which that event may be attributed, the greater the likelihood of that cause {given that event}.

$P(C_i \mid E) \propto P(E \mid C_i)$



"The greater the probability of an observed event given any one of a number of causes to which that event may be attributed, the greater the likelihood of that cause {given that event}. The probability of the existence of any one of these causes {given the event} is thus a fraction whose numerator is the probability of the event given the cause, and whose denominator is the sum of similar probabilities, summed over all causes.

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$$P(C_i | E) = \frac{P(E | C_i) P(C_i)}{P(E)}$$

(Philosophical Essai on Probabilities)

[In general $P(E) = \sum_{j} P(E | C_j) P(C_j)$ (weighted average, with weigths being the probabilities of the conditions) if C_j form a complete class of hypotheses]

$$P(C_i | E) = \frac{P(E | C_i) P(C_i)}{P(E)} = \frac{P(E | C_i) P(C_i)}{\sum_j P(E | C_j) P(C_j)}$$

"This is the fundamental principle ^(*) of that branch of the analysis of chance that consists of reasoning a posteriori from events to causes"

(*) In his "Philosophical essay" Laplace calls 'principles' the 'fundamental rules'.



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Most convenient way to remember Bayes theorem

Laplace's teaching

$\frac{P(H_0 \mid \text{data})}{P(H_1 \mid \text{data})} = \frac{P(\text{data} \mid H_0)}{P(\text{data} \mid H_1)} \times \frac{P(H_0)}{P(H_1)}$

• We should possibly use the <u>data</u>, rather then the test variables ' θ ' (χ^2 etc);

[although in some case 'sufficient summaries' do exist]



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- If P(data | H_i) = 0, it follows P(H_i | data) = 0:
 ⇒ falsification (the 'serious' one) is a corollary of the theorem, rather than a principle.

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- At least two hypotheses are needed!
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- If P(data | H_i) = 0, it follows P(H_i | data) = 0:
 ⇒ falsification (the 'serious' one) is a corollary of the theorem, rather than a principle.
- ▶ There is no conceptual problem with the fact that $P(\text{data} | H_1) \rightarrow 0$ (e.g. 10^{-37}), provided the ratio $P(\text{data} | H_0)/P(\text{data} | H_1)$ is not undefined.

Bayes factor ('likelihood ratio')

$$\frac{P(H_0 \mid \text{data})}{P(H_1 \mid \text{data})} = \frac{P(\text{data} \mid H_0)}{P(\text{data} \mid H_1)} \times \frac{P(H_0)}{P(H_1)}$$



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Prob. $ratio|_{posterior}$ = Bayes factor × Prob. $ratio|_{prior}$

(prior/posterior w.r.t. data)



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Prob. ratio $|_{posterior}$ = Bayes factor × Prob. ratio $|_{prior}$ (prior/posterior w.r.t. data)

If H_0 and H_1 are 'complementary', that is $H_1 = \overline{H}_0$, then

posterior odds = Bayes factor \times prior odds

Telling it with Gauss' words

A quote from the Princeps Mathematicorum (Prince of Mathematicians) is <u>a must</u>.



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Arguments used to derive Gaussian distribution

- $f(\mu | \{x\}) \propto f(\{x\} | \mu) \cdot f_0(\mu)$
- $f_0(\mu)$ 'flat' (all values a priory equally possible)
- posterior maximized at $\mu = \overline{x}$

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<u>Note</u>: indeed Gauss had also invented the "Bayes Factor"! (GdA, arXiv:2003.10878 [math.HO]) Probabilistic inference/prediction

applied to the 'binomial' case

Namely the original problem tackled by Laplace and Bayes



Probabilistic inference/prediction

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Namely the original problem tackled by Laplace and Bayes, but in modern notation and making use of a graphical model:



Probabilistic inference/prediction

applied to the 'binomial' case

Namely the original problem tackled by Laplace and Bayes, but in modern notation and making use of a graphical model:

- 1. draw the graphical model;
- 2. write down the joint pdf of all variables entering the game;
- use Bayes theorem in order to condition on what is known/assumed;
- 4. marginalize over all variables on which we are not interesting;
- 5. do somehow the math.

General case



General case

Model







$$f(x, p, n) = f(x | p, n) \cdot f(p, n)$$





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= $f(x | p, n) \cdot f(p | n) \cdot f(n)$





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(*n* and *p* are independent)

n independent Bernoulli processes Usual case $\rightarrow n$ fixed (for the moment)

Model





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Model



Joint pdf

$f(x,p \mid n) = f(x \mid p,n) \cdot f(p)$



Graphical models of the typical problems





Graphical models of the typical problems





 $\rightarrow f(p \mid n, x)$

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$$f(p | x, n) = \frac{f(p, x | n)}{f(x | n)}$$
$$= \frac{f(x | n, p) \cdot f_0(p)}{f(x | n)}$$

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$$\propto f(x|n,p) \cdot f_0(p)$$
(denominator just normalization)

)

We just need to make explicit f(x | n, p):

$$f(x \mid n, p) = \binom{n}{x} p^{x} (1-p)^{n-x}$$



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$$f(p \mid x, n) = \frac{\frac{n!}{(n-x)! \cdot x!} p^{\times} (1-p)^{n-x} f_{\circ}(p)}{\int_{0}^{1} \frac{n!}{(n-x)! \cdot x!} p^{\times} (1-p)^{n-x} f_{\circ}(p) dp}$$

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(The binomial coefficient is irrelevant, not depending on p)

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For teaching purposes we start from a uniform prior, i.e. $f_o(p) = 1$:

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- The integral at the denominator is the special function "β" (also defined for real values of x and n).
- In our case these two numbers are integer and the integral becomes equal to

$$\frac{x!(n-x)!}{(n+1)!}$$

Solution for uniform prior

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Summaries of the posterior distribution

$$p_m = mode(p) = \frac{x}{n}$$



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Summaries of the **posterior distribution**

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$$E(p) = \frac{x+1}{n+2}$$
"recursive Laplace formula"
("Laplace's rule of succession")
$$Var(p) = \frac{(x+1)(n-x+1)}{(n+3)(n+2)^2}$$

$$= \frac{x+1}{n+2} \left(\frac{n+2}{n+2} - \frac{x+1}{n+2}\right) \frac{1}{n+3}$$

$$= E(p) (1 - E(p)) \frac{1}{n+3}$$

About the meaning of E(p)

- We have used the "first" (*) n trials to learn about "p".
 - [^(*) "First" does not imply time order, but just order in usage.]



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E(p) (and not the mode!) is the probability of every 'future' event which is believed to have the same p of the 'previous' ones.

Related recent applications

(with many details, including general introduction to the relevant ideas and methods, and program samples)

 GdA and A. Esposito, Checking individuals and sampling populations with imperfect tests, (arXiv:2009.04843 [q-bio.PE])

 GdA and A. Esposito, What is the probability that a vaccinated person is shielded from Covid-19? A Bayesian MCMC based reanalysis of published data with emphasis on what should be reported as 'efficacy', (arXiv:2102.11022 [stat.AP])

Large number behaviour: summary

When

- *n* large;
- ► x large;



Large number behaviour: summary

When

- n large;
- ► x large;
- ▶ and (n x) large



Large number behaviour: summary

When

- n large;
- ► x large;
- and (n-x) large

(remember: in the binomial what is 'success' and what is 'failure'

is not absolute: $p \longleftrightarrow q = 1 - p$)



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$$\begin{split} \mathsf{E}(p) &\approx \frac{x}{n} \\ \sigma(p) &\approx \frac{1}{\sqrt{n}} \sqrt{\frac{x}{n} \left(1 - \frac{x}{n}\right)} \end{split}$$

— f(p | x, n) tends to Gaussian



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- $f(p \mid x, n)$ tends to Gaussian, a reflection of the Gaussian limit of $f(x \mid p, n)$
- The probability of a future events is evaluated from the relative frequency of the past events
- No need of 'frequentistic definition' !

Mathematically convenient priors

Before the advent of powerful computers, applying Laplace' ideas ("Bayesian") has always been a severe problem!



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 \rightarrow Computational barrier



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Inferring the "Bernoulli's p"

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 (In particular, the Gaussian is self-conjugate, which is not so great...)

Predicting future nr. of successes and future frequences

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- But we are not sure about it: we need to take into account all possible values, each weighted by f(p)





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- More precisely,

$$f(x_1 \mid n_1, n_0, x_0) = \int_0^1 f(x_1 \mid n_1, p) f(p \mid x_0, n_0) \, dp$$

 $\blacktriangleright X_1 \to f_1$

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X₁ → f₁ (Predicting a future frequency from a past frequency)

Some examples

$f(x_1 \mid n_0, x_0, n_1 = 10)$ in %					
<i>X</i> ₁	$\frac{X_1}{n_1}$	$\int x_0 = 1$	$\int x_0 = 10$	$\int x_0 = 100$	$\int x_0 = 1000$
		$\int n_0 = 2$	$\int n_0 = 20$	$\int n_0 = 200$	$\int n_0 = 2000$
0	0	3.85	0.42	0.12	0.10
1	0.1	6.99	2.29	1.11	0.99
2	0.2	9.44	6.51	4.67	4.42
3	0.3	11.19	12.54	11.88	11.74
4	0.4	12.24	18.07	20.21	20.48
5	0.5	12.59	20.33	24.02	24.55
6	0.6	12.24	18.07	20.21	20.48
7	0.7	11.19	12.54	11.88	11.74
8	0.8	9.44	6.51	4.67	4.42
9	0.9	6.99	2.29	1.11	0.99
10	1	3.84	0.42	0.12	0.10
$E(X_1)$		5	5	5	5
$\sigma[X_1]$		2.64	1.87	1.62	1.58

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$$\propto f(p, x_1, n_0, x_0, n_1)$$

- \Rightarrow The denominator is just a constant.
- ⇒ Very important observation in order to solve the problem numerically or by Monte Carlo methods! (And remember that the numerator can be obtained using the chain rule)

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Terminology:

- nodes (observed/unobserved);
- child/childred;
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Software to analyse it:

- instructions which remind the description of the model by a suitable chain rule;
- computation performed by Markov Chain Monte Carlo.

Joint inference and prediction in JAGS

JAGS: Just Another Gibbs Sampler

(The Gibbs Sampler is an MCMC algorithm,

but the software also uses Metropolis in though cases)





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Model:

```
model{
    x0 ~ dbin(p, n0)
    x1 ~ dbin(p, n1)
    p ~ dbeta(1, 1) # flat prior in termes of a Beta
}
```

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Use of JAGS from R via rjags (n0 = 20, x0 = 10, n1 = 10)





Trace of x1



Density of x1



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Use of JAGS from R via rjags (n0 = 20, x0 = 10, n1 = 10)











(10000 samples).

 $p = 0.498 \pm 0.105; x_1 = 4.98 \pm 1.86$

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Use of JAGS from R via rjags

Scatter plot of sampled $f(p, x_1 | n_0, x_0, n_1)$



cor(p,x1) = 0.56

x1

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A package more suited for Physics analysis (expecially HEP)



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Presently rewritten in Julia: https://github.com/bat/BAT.jl

n independent Bernoulli processes





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- we have recorded x 'signals';
- how many particles impinged the detector? $\longrightarrow f(n | x, p)$?

Extending the model

Our problem (but in Physics it is often not so simple)





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But we need some (usually indirect) knowledge about p



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But what is *n*?

In Physics we are usually not interested in the numbers we do see, but in those which have 'physical meaning'.



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But λ is not really physical $\longrightarrow \lambda = r T$ (*r*: intensity of the Poisson process).

 $\lambda = r \cdot T$:



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(Dashed arrows used in literature for deterministic links)

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(Dashed arrows used in literature for deterministic links) In JAGS, e.g., lambda <- r * T;</pre>

Remembering that *p* was got from a measurement:



The rate r gets contributions from signal and background



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But, since $r = r_S + r_B$,

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(T_0 and T assumed to be measured with sufficient accuracy)





(*) Assuming unity efficiency



All the rest is a technical question of

- writing down the joint pdf of all variables;
- (re-)conditioning on the assumed/observed quantities;
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Or, more easily, use software **grounded on probability theory**, like BUGS, JAGS, BAT, etc.

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- Gaussian model;
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More detailed applications (including scripts) in

https://www.roma1.infn.it/~dagos/prob+stat.html
[For BAT and BAT.jl see their web pages]

The probabilistic framework basically set up by Laplace^(*) in his monumental work is healthy and grows up well.

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 - you are very close to the solution, although the analitic one is often out of reach (MCMC's rescue us!)

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Handling systematics in the probabilistic approach

(Answering to a question: diagrams show the case of uncertain offset systematics, best known to be, after proper calibration, $z = 0 \pm \sigma_z$)



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References

As starting points, particularly recommended are the last papers, in which graphical models are systematically exploited:

- arXiv:2001.03466 [physics.data-an]
- arXiv:2009.04843 [q-bio.PE]
- arXiv:2012.04455 [stat.ME]
- arXiv:2102.11022 [stat.AP]
- Much more can be found in
 - https://www.roma1.infn.it/~dagos/prob+stat.html
 - https://www.roma1.infn.it/~dagos/dott-prob_31/